

Метод зеркального спуска в задачах о многоруком бандите

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На основе совместных работ с

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План

1. Краткое введение. Идея МЗС (в непрерывном времени) и некоторые его свойства.

Роль преобразования Лежандра, функция Ляпунова, усреднение траектории исходного пространства (в оптимизации).

Оценка скорости сходимости по оптимизируемой функции. Некоторые выводы.

План (продолжение)

2. Общие понятия, объекты и конструкции: исходная и двойственная норма, прокси-функция на заданном выпуклом компакте и ее сопряженная (преобразование Лежандра-Фенхеля), их свойства (при условии сильной выпуклости).

Примеры: “евклидовы” случаи как во всем пространстве, так и в шаре, и энтропийная прокси-функция на стандартном симплексе и распределение Гиббса.

3. Приложение МЗС к задаче о многоруком бандите.

4. Краткий список литературы.

Introduction

Mirror Descent Method (MDA) is a gradient-type recursive method for convex optimization, i.e. primal-dual method performing the descent in a dual space and mapping the resulted points to a primal space. See the following references:

1. Nemirovski and Yudin (1979/1983): [1]
2. Ben-Tal, Margalit, and Nemirovski (2001): [2]
3. Beck and Teboulle (2003): [3]
4. Nesterov (2005, 2007): [4], [5]
5. Juditsky, Nazin, Tsybakov, and Vayatis (2005): [6]
6. Juditsky, Lan, Nemirovski, and Shapiro (2007): [7]

1 Idea behind MDM (continuous time) [1]

Consider a primal-dual method, that is MDM:

$$\dot{\xi}(t) = -\nabla_x f(x(t)), \quad \xi(0) = \xi_0, \quad (1)$$

$$x(t) = \nabla_{\xi} W(\xi(t)), \quad t \geq 0. \quad (2)$$

Here:

- f is a convex function to be minimized in Banach space E ,
- W is a uniform differentiable, convex function on dual space E^* .

As an example, “Euclidean” case of

$$W(\xi) = \frac{1}{2} \|\xi\|_2^2$$

gives a well-known standard gradient method

$$\dot{x}(t) = -\nabla_x f(x(t)).$$

Let us look at a simple analysis as follows.

Assume

$$x^* = \arg \min f(x).$$

Then we have a candidate Lyapunov function

$$W_*(\xi) \triangleq W(\xi) - \langle \xi, x^* \rangle,$$

since

$$\frac{dW_*(\xi(t))}{dt} = \langle \dot{\xi}(t), \nabla_{\xi} W(\xi(t)) - x^* \rangle \quad (3)$$

$$= - \langle \nabla_x f(x(t)), x(t) - x^* \rangle \quad (4)$$

$$\leq f(x^*) - f(x(t)) \quad (5)$$

$$\leq 0, \quad (6)$$

that is function $W_*(\xi)$ decreases along the trajectory $\{\xi(t)\}$.

Furthermore, (3)–(5) lead to

$$f(x(t)) - f(x^*) \leq \langle \dot{\xi}(t), x^* \rangle - \frac{dW(\xi(t))}{dt}, \quad (7)$$

and, assuming that

$$\xi(0) = 0, \quad W(0) = 0,$$

and integrating by $t \in [0, T]$, we get

$$\begin{aligned} \int_0^T f(x(t))dt - T f(x^*) &\leq \langle \xi(T), x^* \rangle - W(\xi(T)) \quad (8) \\ &\leq V(x^*) \quad (9) \end{aligned}$$

with the Legendre transformation

$$V(x) \triangleq \sup_{\xi} \{ \langle \xi, x \rangle - W(\xi) \}.$$

Now, introduce the average estimate

$$\hat{x}(T) \triangleq \frac{1}{T} \int_0^T x(t) dt .$$

By Jensen's inequality, due to convexity of $f(x)$, eqs (8)–(9) lead to

$$f(\hat{x}(T)) - f(x^*) \leq \frac{1}{T} V(x^*). \quad (10)$$

Remark: The rate $O(1/T)$ in the upper bound above changes for that of $O(1/\sqrt{T})$ when working with discrete time gradient observations.

Résumé:

- Function $W : E^* \rightarrow \mathbb{R}$ is a parameter of MDM which ensures the Lyapunov function $W_* : E^* \rightarrow \mathbb{R}$; in particular, MDM reduces to standard gradient method; therefore, this additional degree of freedom may improve the accuracy algorithm, at least potentially.
- MDM leads to the average estimate $\hat{x}(t)$, i.e. time-average to current estimates over the time interval $[0, t]$.

- Non-asymptotical upper bound on difference between current estimation function $f(\hat{x}(t))$ and function minimum $f(x^*)$ is ensured; this upper bound is of type $O(T^{-1})$, and it is directly depending on $V(x^*)$; therefore, the given class function has to ensure the finite upper bound $\sup V(x)$. (Thus, further consideration is reduced to function minimization over a given compact convex set.)
- The previous consideration shows the role of Legendre transformation.

2 A Generalized View-Point

Proxy functions. Denote by E the space \mathbb{R}^M with a norm $\|z\|$ and by E^* the dual space which is \mathbb{R}^M equipped with the conjugate (dual) norm

$$\|z\|_* = \max_{\|\theta\|=1} z^T \theta, \quad \forall z \in E^*.$$

Let Θ be a convex, closed set in E . For a given parameter $\beta > 0$ and a convex function $V : \Theta \rightarrow \mathbb{R}$, we call β -conjugate function of V the Legendre–Fenchel type transform of βV :

$$\forall z \in E^*, \quad W_\beta(z) = \sup_{\theta \in \Theta} \{-z^T \theta - \beta V(\theta)\}. \quad (11)$$

Assumption (L). *A convex function $V : \Theta \rightarrow \mathbb{R}$ is such that its β -conjugate W_β is continuously differentiable on E^* and its gradient ∇W_β satisfies*

$$\|\nabla W_\beta(z) - \nabla W_\beta(\tilde{z})\| \leq \frac{1}{\alpha\beta} \|z - \tilde{z}\|_*, \quad \forall z, \tilde{z} \in E^*, \beta > 0,$$

where $\alpha > 0$ is a constant independent of β .

Assumption (L) relates to the strong convexity w.r.t. *initial norm* $\| \cdot \|$:

$$V(sx + (1 - s)y) \leq sV(x) + (1 - s)V(y) - \frac{\alpha}{2}s(1 - s)\|x - y\|^2 \quad (12)$$

for all $x, y \in \Theta$ and any $s \in [0, 1]$.

The following proposition sums up some properties of β -conjugates and, in particular, yields a sufficient condition for Assumption (L).

Proposition 1. *Let function $V : \Theta \rightarrow \mathbb{R}$ be convex and $\beta > 0$. Then, the β -conjugate W_β of V has the following properties.*

1. *The function $W_\beta : E^* \rightarrow \mathbb{R}$ is convex and has a conjugate βV , i.e.,*

$$\forall \theta \in \Theta, \quad \beta V(\theta) = \sup_{z \in E^*} \{ -z^T \theta - W_\beta(z) \} .$$

2. *If function V is α -strongly convex with respect to the initial norm $\| \cdot \|$ then*

- (i) *Assumption (L) holds true,*
- (ii) $\operatorname{argmax}_{\theta \in \Theta} \{ -z^T \theta - \beta V(\theta) \} = -\nabla W_\beta(z) \in \Theta .$

Definition 1. We call $V : \Theta \rightarrow \mathbb{R}_+$ proxy function if it is convex, and

(i) there exists a point $\theta_* \in \Theta$ such that $\min_{\theta \in \Theta} V(\theta) = V(\theta_*)$,

(ii) Assumption (L) holds true.

Example 1: Consider Euclidean space \mathbb{R}^M as set $\Theta = \mathbb{R}^M$. Then half of the squared Euclidean norm be related proxy-function

$$V(\theta) = \frac{1}{2} \|\theta\|^2, \quad \theta \in \mathbb{R}^M.$$

Indeed, minimum point $\theta_* = 0 \in \mathbb{R}^M$, the function is strongly convex w.r.t. the Euclidean norm, and the constant of strong convexity $\alpha = 1$. Evidently, $E^* = E$, a β -conjugate function

$$W_\beta(z) = \frac{1}{2\beta} \|z\|^2, \quad z \in \mathbb{R}^M$$

with $\nabla W_\beta(z) = z/\beta$. □

Example 2: Let set Θ in the previous Example be Euclidean r -ball with the center at the origin, $r > 0$. The same proxy-function leads to the related β -conjugate function as follows: $\forall z \in \mathbb{R}^M$,

$$W_{\beta}(z) = \begin{cases} \frac{1}{2\beta} \|z\|^2, & \|z\| \leq r\beta, \\ r\|z\| - \frac{\beta}{2} r^2, & \text{otherwise.} \end{cases}$$

The gradient

$$\nabla W_{\beta}(z) = \begin{cases} \frac{1}{\beta} z, & \|z\| \leq r\beta, \\ rz/\|z\|, & \text{otherwise;} \end{cases}$$

it realizes the metric projection onto ball $B_{r\beta}$. □

Example 3: Consider a standard simplex $\Theta = \Theta_M$ and an entropy-type proxy function

$$V(\theta) = \ln(M) + \sum_{j=1}^M \theta^{(j)} \ln \theta^{(j)} \quad (13)$$

(where $0 \ln 0 \triangleq 0$) which has a single minimizer $\theta_* = (1/M, \dots, 1/M)^T$ with $V(\theta_*) = 0$.

Let the initial norm in \mathbb{R}^M be 1-norm

$$\|\theta\|_1 = \sum_{j=1}^M |\theta^{(j)}|, \quad \theta \in \mathbb{R}^M.$$

Therefore, the initial space is $E = \ell_1^M$, and the dual space $E^* = \ell_\infty^M$ is \mathbb{R}^M equipped with the sup-norm

$$\|z\|_\infty = \max_{\|\theta\|_1=1} z^T \theta = \max_{1 \leq j \leq M} |z^{(j)}|, \quad \forall z \in E^*.$$

It is directly checked that this function is α -strongly convex w.r.t. the 1-norm, with the parameter

$$\alpha = 1.$$

This leads to β -conjugate function to $V(\theta)$ as follows:

$$W_\beta(z) = \beta \ln \left(\frac{1}{M} \sum_{k=1}^M e^{-z^{(k)}/\beta} \right), \quad z \in \mathbb{R}^M, \quad (14)$$

with partial derivatives relating to a Gibbs distribution on the coordinates of vector $z = (z^{(1)}, \dots, z^{(M)})^T$, with β being a “temperature” parameter:

$$-\frac{\partial W_\beta(z)}{\partial z^{(j)}} = e^{-z^{(j)}/\beta} \left(\sum_{k=1}^M e^{-z^{(k)}/\beta} \right)^{-1}, \quad j = 1, \dots, M. \quad (15)$$

□

Convex Stochastic Optimization Problem

$$A(\theta) \triangleq \mathbb{E} Q(\theta, Z) \rightarrow \min_{\theta \in \Theta}$$

with loss function $Q : \Theta \times \mathcal{Z} \rightarrow \mathbb{R}_+$ being such that the random function $Q(\cdot, Z) : \Theta \rightarrow \mathbb{R}_+$ is convex a.s., on a convex closed set $\Theta \subset \mathbb{R}^M$.

Let a learning sample be given in the form of an i.i.d. sequence (Z_1, \dots, Z_{t-1}) , where each Z_i has the same distribution as Z .

Denote stochastic subgradients

$$u_i(\theta) = \nabla_{\theta} Q(\theta, Z_i), \quad i = 1, 2, \dots, \quad (16)$$

which are measurable functions on $\Theta \times \mathcal{Z}$ such that, for any $\theta \in \Theta$, the expectation $\mathbb{E} u_i(\theta)$ belongs to the subdifferential of the function $A(\theta)$.

Mirror Descent Algorithm (MDA)

The algorithm is defined as follows:

- Fix the initial value $\zeta_0 = 0 \in \mathbb{R}^M$.
- For $i = 1, \dots, t - 1$, do the recursive update

$$\begin{aligned}\zeta_i &= \zeta_{i-1} + \gamma_i u_i(\theta_{i-1}), \\ \theta_i &= -\nabla W_{\beta_i}(\zeta_i).\end{aligned}\tag{17}$$

- Output at iteration t the following convex combination:

$$\hat{\theta}_t = \sum_{i=1}^t \gamma_i \theta_{i-1} \left(\sum_{i=1}^t \gamma_i \right)^{-1}.\tag{18}$$

3 Multi-Armed Bandit Problem (classic).

Presented at the 17th IFAC World Congress:

- Juditsky, A., A.V. Nazin, A.B. Tsybakov, N. Vayatis.
Gap-free Bounds for Stochastic Multi-Armed Bandit.
Proc. 17th IFAC World Congress, Seoul, Korea, 6–11 July 2008, pp.11560–11563.

Let $X = \{x(1), \dots, x(N)\}$ be a set of N available actions. At each time $t = 1, 2, \dots$, we have to choose sequentially an action $x_t \in X$. We denote by η_t the observable (instantaneous) loss for the choice of x_t , and introduce the average loss up to horizon T which is to be minimized:

$$\Phi_T = \frac{1}{T} \sum_{t=1}^T \eta_t. \quad (19)$$

A strategy \mathcal{U} is a sequence of rules for the choice x_t at times $t = 1, \dots, T$. In the stochastic setup that we consider here, the sequence of losses $(\eta_t)_{t \geq 1}$ is a stochastic process and x_t is a measurable function (random, in general) depending only on the vector of past decisions and losses

$$(x_1, \dots, x_{t-1}; \eta_1, \dots, \eta_{t-1}).$$

Any strategy \mathcal{U} generates a flow of σ -algebras

$\mathcal{F}_t = \sigma\{x_1, \dots, x_t; \eta_1, \dots, \eta_t\}$, $t \geq 1$ (for brevity we do not indicate the dependence of \mathcal{F}_t on \mathcal{U}). Throughout the paper we denote by $z^{(j)}$ the j th component of vector $z \in \mathbb{R}^N$.

Two basic assumptions:

A1. With probability 1, the conditional expectations satisfy

$$\mathbb{E}\{\eta_t \mid \mathcal{F}_{t-1}, x_t = x(k)\} = a_k, \quad k = 1, \dots, N, \quad (20)$$

where $a_k \in \mathbb{R}$ are unknown deterministic values.

The value a_k characterizes the expected loss for deciding to take the action $x_t = x(k)$ at time t . Assumption A1 says that this loss should not depend on t .

A2. The second conditional moment of the loss η_t is a.s. bounded by a constant:

$$\mathbb{E}\{\eta_t^2 \mid \mathcal{F}_{t-1}, x_t\} \leq \sigma^2 < \infty. \quad (21)$$

It is easy to prove (see, e.g., [8]) that under these assumptions all the limiting points of the average loss sequence $(\Phi_t)_{t \geq 1}$ cannot be almost surely (a.s.) less than

$$a_{\min} \triangleq \min_{k=1, \dots, N} a_k .$$

Thus, the problem is to design a strategy \mathcal{U}^* which has the asymptotically minimal average loss:

$$\Phi_T \rightarrow a_{\min} \quad \text{as} \quad T \rightarrow \infty , \quad (22)$$

in an appropriate probability sense.

We study here *convergence in mean*, trying to get the rate of convergence

$$\mathbb{E}(\Phi_T) \rightarrow a_{\min}$$

as fast as possible.

In particular, we provide *non-asymptotic* upper bounds for the expected excess risk $\mathbb{E}(\Phi_T) - a_{\min}$ that are close, up to logarithmic factors, to the lower bound of the order $\sqrt{N/T}$ proved for arbitrary N by (see Theorem 6.11 in [10]).

We will suppose that the following assumption on the loss sequence $(\eta_t)_{t \geq 1}$ holds:

A3. The losses are nonnegative: $\eta_t \geq 0$ a.s.

Below we propose a randomized decision strategy in which, at each step $t + 1$, the action x_{t+1} is drawn according to a distribution $p_t \triangleq \left(p_t^{(1)}, \dots, p_t^{(N)} \right)^\top$ over X where:

$$p_t^{(k)} \triangleq \mathbb{P}(x_{t+1} = x(k) \mid \mathcal{F}_t), \quad k = 1, \dots, N. \quad (23)$$

The update of the distribution p_t over time is given by the MDA.

Denote by Θ the simplex of all probability vectors over X :

$$\Theta \triangleq \left\{ p \in \mathbb{R}_+^N \mid \sum_{k=1}^N p^{(k)} = 1 \right\} . \quad (24)$$

We then define the mean (over the set of actions) loss function A on Θ :

$$A(p) = \sum_{k=1}^N a_k p^{(k)} = a^\top p, \quad p \in \Theta, \quad (25)$$

where $a = (a_1, \dots, a_N)^\top$. Since p_t is a random vector, the quantity $A(p_t)$ is a random variable. The update rule for the probability distribution p_t uses a stochastic gradient of A .

The expected average loss equals to the average over time of the expectations $\mathbb{E}A(p_t)$, that is

$$\mathbb{E}(\Phi_T) = \frac{1}{T} \sum_{t=1}^T \mathbb{E}(\mathbb{E}(\eta_t | x_t, \mathcal{F}_{t-1})) = \frac{1}{T} \sum_{t=1}^T \mathbb{E}(A(p_{t-1})) . \quad (26)$$

Theorem. *Let assumptions A1–A3 be satisfied and let the conditional distributions $(p_t)_{t \geq 0}$ be defined by the MDA.*

Then, for any horizon $T \geq 1$,

$$\mathbb{E}(\Phi_T) - a_{\min} \leq 2\sigma \frac{\sqrt{(T+1)N \ln N}}{T} . \quad (27)$$

The MD algorithm for multi-armed bandit.

1. Fix $p_0 = (N^{-1}, \dots, N^{-1})^T$ and $\zeta_0 = 0 \in \mathbb{R}^N$.
2. For $t = 1, \dots, T$:
 - (a) draw an action $x_t = x(k_t)$ with random k_t distributed according to p_{t-1} ;
 - (b) compute the stochastic gradient

$$u_t(p_{t-1}) = \frac{\eta_t}{p_{t-1}(k_t)} e_N(k_t); \quad (28)$$

- (c) update the dual and probability vectors

$$\zeta_t = \zeta_{t-1} + \gamma_t u_t(p_{t-1}), \quad (29)$$

$$p_t = -\nabla W_{\beta_t}(\zeta_t). \quad (30)$$

3. At horizon $t = T$, output a sequence of actions (x_1, \dots, x_T) .

The tuning parameters γ_t and β_t are as follows: $\forall t \geq 1$,

$$\gamma_t \equiv 1, \quad \beta_{t-1} = \beta_0 \sqrt{t}, \quad \beta_0 = \sigma \sqrt{N / (\ln N)}. \quad (31)$$

Notice that

$$\mathbb{E} \left\{ \frac{\eta_t}{p_{t-1}^{(k_t)}} e_N(k_t) \middle| \mathcal{F}_{t-1} \right\} = a = \nabla A(p_{t-1}). \quad (32)$$

Here is the corrected information lower bound from [10], Theorem 6.11. Let $\Phi_{i,T}$ be mean losses under fixed i -th arm, i.e., $x_t \equiv x(i)$.

Theorem. *Let $T, N \geq 1$ be such that $T > N/(4 \ln(4/3))$. There exists a loss function such that for any, possibly randomized, control strategy*

$$\sup_{\mathcal{Z}} \mathbb{E} \left(\Phi_T - \min_{i=1, \dots, N} \Phi_{i,T} \right) \geq \frac{\sqrt{N/T}}{32 \sqrt{\ln 4/3}}, \quad (33)$$

where sup is over set of all multi-armed bandit problems with losses η_t with values from interval $[0, 1]$ a.s.

Remarks: The information lower bound above (see [10], Theorem 6.11) differs from the upper bound (27) by logarithmic term $\sqrt{\ln N}$.

The wrong constant in the lower bound of Theorem 6.11 [10]

$$\frac{\sqrt{2} - 1}{\sqrt{32 \ln 4/3}} \approx 0.1365 \quad (34)$$

is more than that of Theorem

$$\frac{1}{32\sqrt{\ln 4/3}} \approx 0.0583. \quad (35)$$

Unfortunately, the constant (34) is uncorrectly calculated in [10], page 165.

4 Multi-Armed Bandit Governed by a Stationary Finite Markov Chain.

To be presented at the ECC2013:

- Nazin, A.V., B.M. Miller. Mirror Decent Algorithm for a Multi-Armed Bandit Governed by a Stationary Finite State Markov Chain. *The 12th European Control Conference, ECC13, July 17–19, 2013, Zurich, Switzerland.*

In addition to the classic case of Multi-Armed Bandit Problem, assume that instantaneous losses η_t depend now on both chosen arm $x_t \in X$ and current state $z_t \in Z$ of *unknown* stationary finite Markov Chain (MC), $Z = \{z(1), \dots, z(K)\}$. The main new assumption is as follows:

- the transition probabilities of the state $z_t \in Z$ at each time $t \in \{0, 1, \dots\}$ to the next state $z_{t+1} \in Z$ are presented by unknown conditional probabilities: $\forall t$,

$$\mathbb{P}\{z_{t+1} = z(j) \mid z_t = z(i)\} = \pi_{ij}; \quad (36)$$

- MC state z_t is observable at current time $t \geq 0$.

Further assumptions:

A1. For each $t = 1, 2, \dots$ the sets of random variables

$$\{\eta_t(z, u, \omega) \mid z \in Z, u \in U\} \quad \text{and}$$

$$\{\eta_s(z, u, \omega), z_k, u_k \mid z \in Z, u \in U, s = \overline{1, t-1}, k = \overline{1, t}\}$$

are independent.

A2. For each $z(i) \in Z$, $u(\ell) \in U$, and $t = 1, 2, \dots$ the losses $\eta_t(z(i), u(\ell), \omega)$ are non-negative a.s. and their *a priori unknown* expectations are time-invariant:

$$\mathbb{E}\{\eta_t(z(i), u(\ell), \omega)\} \triangleq a_{i\ell} \quad \forall t. \quad (37)$$

A3. The losses $\eta_t(z(i), u(\ell), \omega)$ are bounded in the mean

square sense, i.e.

$$\mathbb{E}\{\eta_t^2(z(i), u(\ell), \omega)\} \leq \sigma^2 < \infty. \quad (38)$$

- A4. The Markov chain is regular, i.e., the transition probability matrix Π is regular (i.e., the state set Z represents a unique ergodic class).
- A5. The initial distribution of MC assumed to be stationary. The stationary distribution of the MC states is assumed to be unknown.

Introduce randomized strategy by

$$d_t^{(i\ell)} \triangleq \mathbb{P}\{u_t = u(\ell) \mid z_t = z(i), \mathcal{F}_{t-1}\}. \quad (39)$$

Under a stationary strategy \mathcal{U}_{St} with $d \triangleq \|d^{(i\ell)}\|$, the loss expectation lead to the loss function

$$\mathbb{E}\{\eta_t\} = \sum_{i=1}^K q_i \sum_{\ell=1}^N a_{i\ell} d^{(i\ell)} \quad (40)$$

$$\triangleq A(d), \quad d \in D, \quad (41)$$

with stationary state probabilities

$$q_i \triangleq \mathbb{P}\{z_t = z(i)\} \quad (42)$$

and the set stochastic matrix

$$D \triangleq \left\{ d \mid d^{(i\ell)} \geq 0, \sum_{\ell=1}^N d^{(i\ell)} = 1 \ (i = \overline{1, K}, \ell = \overline{1, N}) \right\} .$$

Denote

$$A_{\min} \triangleq \min_{d \in D} A(d) . \quad (43)$$

Theorem. *Let assumptions A1–A5 be satisfied and let the conditional distributions $(d_t^{(i)})_{t \geq 0}$, $i = \overline{1, K}$, be defined by the randomized control algorithm (see below) with parameters (48). Then, for any time $T \geq 1$,*

$$\mathbb{E} (\Phi_T) - A_{\min} \leq 2\sigma \sqrt{KN \ln N} \frac{\sqrt{T+1}}{T} . \quad (44)$$

■

Thus, we fix the increasing temperature parameter sequence $(\beta_t)_{t \geq 0}$ and introduce the control randomized strategy as follows.

1. Fix the initial matrix d_0 with equal entries, i.e., $d_0^{(ij)} \equiv 1/N$, and zero dual matrix $\zeta_0 = 0 \in \mathbb{R}^{K \times N}$;

(a) for each $t \geq 0$, by having the observed state $z_t = z(i_t)$, draw arm $x_t = x(\ell_t)$ with random $\ell_t \in \overline{\{1, N\}}$ distributed according to $(d_t^{(i_t 1)}, \dots, d_t^{(i_t N)})^\top$;

(b) compute a stochastic gradient

$$\Xi_{t+1} = \frac{\eta_{t+1}}{d_t^{i_t \ell_t}} e_K(i_t) e_N^\top(\ell_t); \quad (45)$$

(c) update both dual and initial variables

$$\zeta_{t+1} = \zeta_t + \Xi_{t+1}, \quad (46)$$

$$d_{t+1}^{(i)} = G_{\beta_t}(\zeta_{t+1}^{(i)}), \quad \forall i = \overline{1, K}. \quad (47)$$

2. *At time T of interest, output the observed sequences of states (z_0, \dots, z_T) , control actions (u_0, \dots, u_T) , matrices (d_0, \dots, d_T) , and the observed losses $(\eta_1, \dots, \eta_{T+1})$ and Φ_T .*

The tuning algorithm parameter β_t is defined as follows:

$\forall t = 0, 1, \dots,$

$$\beta_t = \beta_0 \sqrt{t + 1}, \quad \beta_0 = \sigma \sqrt{N / (K \ln N)}. \quad (48)$$

Вычислительный пример: $K = 7$ и $N = 5$;

$$\|\pi_{ij}\| = \begin{pmatrix} 1/4 & 1/2 & 0 & 0 & 0 & 0 & 1/4 \\ 1/4 & 1/4 & 1/2 & 0 & 0 & 0 & 0 \\ 0 & 1/4 & 1/4 & 1/2 & 0 & 0 & 0 \\ 0 & 0 & 1/4 & 1/4 & 1/2 & 0 & 0 \\ 0 & 0 & 0 & 1/4 & 1/4 & 1/2 & 0 \\ 0 & 0 & 0 & 0 & 1/4 & 1/4 & 1/2 \\ 1/2 & 0 & 0 & 0 & 0 & 1/4 & 1/4 \end{pmatrix} ;$$

далее, матрица средних потерь

$$\|a_{il}\| = \begin{pmatrix} 0.1 & 0.3 & 0.5 & 0.7 & 0.9 \\ 0.55 & 0.15 & 0.25 & 0.35 & 0.45 \\ 0.325 & 0.375 & 0.175 & 0.225 & 0.275 \\ 0.2375 & 0.2625 & 0.2875 & 0.1875 & 0.2125 \\ 0.175 & 0.225 & 0.325 & 0.375 & 0.275 \\ 0.15 & 0.25 & 0.35 & 0.55 & 0.45 \\ 0.1 & 0.7 & 0.9 & 0.3 & 0.5 \end{pmatrix},$$

и $A_{\min} = 0.1482$. Случайные потери $\eta_t(z(i), x(\ell), \omega)$ в состоянии $z(i)$ и выбранной ручке $x(\ell)$ являются н.о.р. с.в. Бернулли с вероятностями $\mathbb{P}(\eta_t(z(i), x(\ell), \omega) = 1) = a_{il}$.

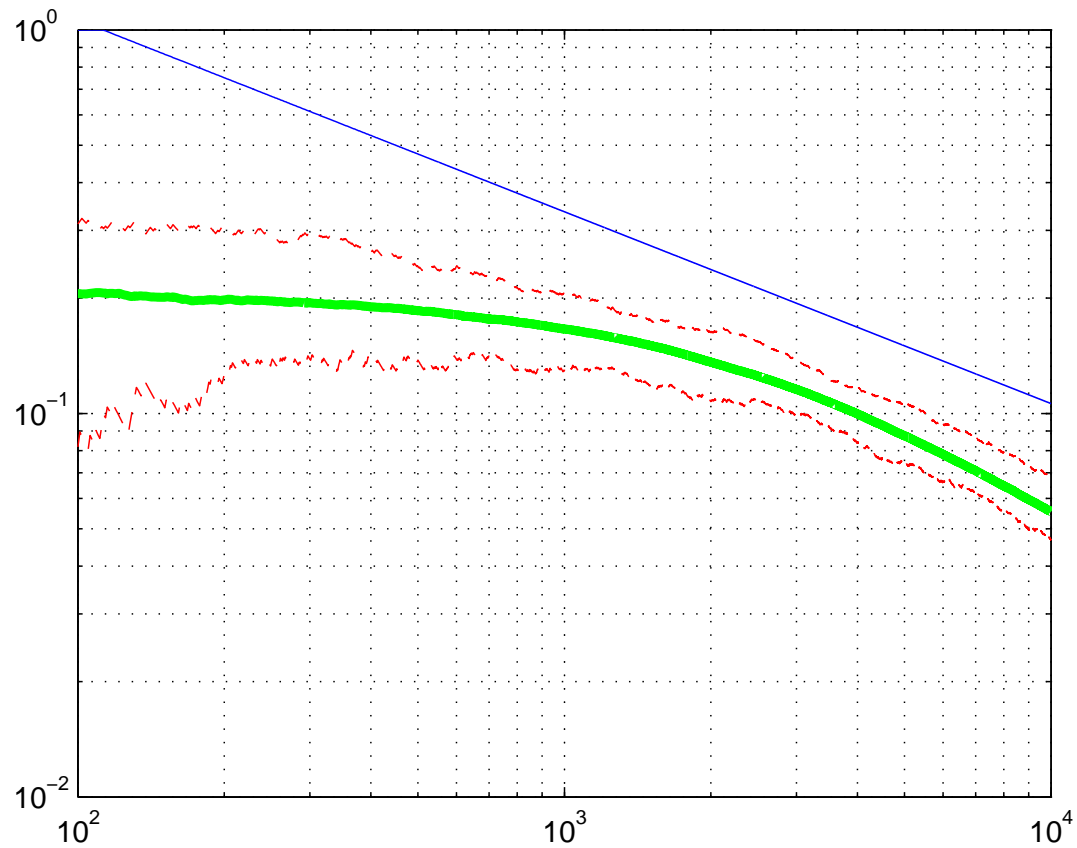


Рис. 1: Результаты вычислительного примера с числом состояний $K = 7$ и числом рук $N = 5$ представлены в двойном логарифмическом масштабе на интервале времени $t = 100, \dots, 10000$.

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THANK YOU FOR YOUR ATTENTION !!!